

# A design-by-treatment interaction model for network meta-analysis with integrated nested Laplace approximations (INLA)

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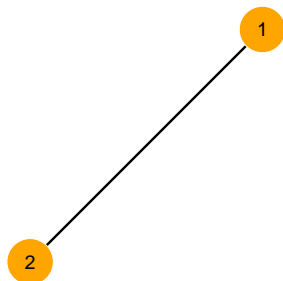
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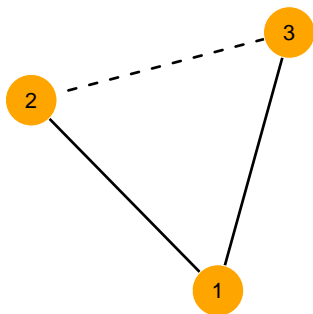


# Pairwise meta-analysis



- Only two treatments are compared
- Trt 1 vs Trt 2 can be **directly** estimated ( $d_{1,2}$ )
- **But**, increasingly, many competing treatments exist
- And multi-arm trials

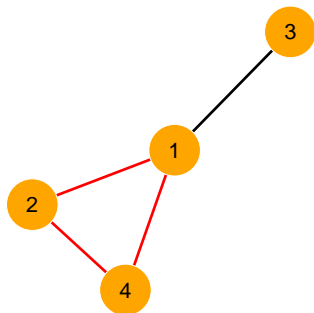
## Need for a broader approach



- Network meta-analysis
- Solid lines indicate comparisons are available
- **Indirect estimate** of 2 vs 3

$$d_{2,3}^{\text{Ind}} = d_{1,2}^{\text{Dir}} - d_{1,3}^{\text{Dir}}$$

# Terminology in NMA (Salanti, 2012)

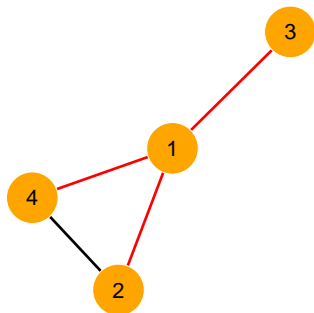


- From Graph theory: vertex, edge
- **Cycle:** Red lines
- **Design:** set of treatments included in a trial; 1-2 design, 1-2-3 design

## Terminology in NMA (cont.)

- **Heterogeneity** in treatment effects between trials  
⇒ As in a pairwise meta-analysis
- **Consistency**: No discrepancy between indirect and direct estimates:  $d_{1,2}^{\text{Dir}} = d_{1,2}^{\text{Ind}}$
- Consistency relation:  $d_{1,2}^{\text{Dir}} = d_{1,3}^{\text{Dir}} - d_{2,3}^{\text{Dir}}$
- Trials of different comparisons were undertaken in different periods
- Right-hand side parameters are **basic parameters** ( $\mathbf{d}_b$ )  
⇒ Parametrization of the network
- Others are **functional parameters** ( $\mathbf{d}_f$ )

# A simple network



- $\mathbf{d}_b = \{d_{12}, d_{13}, d_{14}\}$  (red lines)  
 $\Rightarrow d_f = d_{24} = d_{12} - d_{14}$
- Consistency relation  
 $\Rightarrow$  3-cycle

# Statistical models for NMA

- Hierarchical models, more specifically generalized linear mixed models (GLMMs)
- Contrast-based vs arm-based models
- Trial-arm level instead of summary-level (aggregate-level) approach
  - ⇒ Advantage: the former is one-stage approach
- Datasets with different endpoints (dichotomous, continuous, time-to-event) can be modelled
- Basic model is same, but **likelihood** and **link function** can change

## Consistency models (Dias et al., 2011)

- For convenience, consider data with binomial endpoints
- In trial  $i$ ;  $t_k$  is a treatment arm  
⇒ when  $k = 1$ ,  $t_1$ , is a baseline arm.
- Number of events,  $y_{i,t_k} \sim \text{Bin}(\pi_{i,t_k}, n_{i,t_k})$
- Linear predictor with logit link

$$\text{logit}(\pi_{i,t_k}) = \begin{cases} \mu_i, & \text{if } k = 1 \\ \mu_i + d_{t_1 t_k} + \gamma_{i,t_1 t_k}, & \text{if } k \neq 1. \end{cases}$$

where  $\mu_i$  nuisance parameter and  $d_{t_1 t_k}$  basic parameters

- Heterogeneity random effects:  $\gamma_{i,t_1 t_k} \sim \mathcal{N}(0, \tau^2)$



## Consistency models (cont.)

- But, for a multi-arm trial: dependency within trial!
- Example: A three-arm trial  $i$  with the design 1-2-3
  - $\boldsymbol{\gamma}_i = (\gamma_{i,12}, \gamma_{i,13})^T \sim \mathcal{N}_2(\mathbf{0}, \boldsymbol{\Sigma}_\gamma)$
  - A simple but a convenient structure is as follows (Higgins and Whitehead, 1996):

$$\boldsymbol{\Sigma}_\gamma = \begin{bmatrix} \tau^2 & \tau^2/2 \\ \tau^2/2 & \tau^2 \end{bmatrix}$$

- Models are needed to account for **inconsistency** in the network

## Design-by-treatment interaction model (Higgins et al., 2012)

- **Design inconsistency:** occurs between trials involving different designs
- 1,2,3 trials can be inconsistent with 1,2 trials
- Adding design-specific inconsistency parameters to the consistency model
- Improvement of cycle-inconsistency approach (Lu and Ades, 2006)

# Jackson Model (Jackson et al., 2014)

- Inconsistency parameters as random effects

$$\text{logit}(\pi_{i,t_k}) = \begin{cases} \mu_i, & \text{if } k = 1 \\ \mu_i + d_{t_1 t_k} + \gamma_{i,t_1 t_k} + \omega_{t_1 t_k}^{D(i)}, & \text{if } k \neq 1. \end{cases}$$

$\omega^{D(i)} \sim \mathcal{N}_{T-1}(\mathbf{0}, \Sigma_\omega)$  such that  $\Sigma_\omega$  has diagonal entries  $\kappa^2$  and all others are  $\kappa^2/2$

- NMA-regression: incorporating trial-specific covariates to the model in order to explain sources of heterogeneity and/or inconsistency

# Fully-Bayesian inference for NMA models

## Markov chain Monte Carlo (MCMC)

- A simulation-based technique and the most popular among NMA-analyzers
- Computationally intensive & convergence diagnostics

## Integrated nested Laplace approximations (INLA)

- An approximate Bayesian method (Rue et al., 2009) for latent Gaussian models (LGMs)
- Fast and accurate alternative to MCMC
- **Laplace approximations & numerical integration**
- Implemented in R-INLA (<http://www.r-inla.org/>)

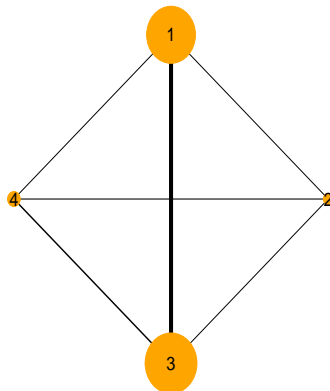
# INLA for NMA models

- How NMA models are LGMs? Three stages:
  - 1 Observational model:  $p(\mathbf{y}|\boldsymbol{\alpha}, \boldsymbol{\Psi})$  where  $\boldsymbol{\alpha} = (\boldsymbol{\mu}, \mathbf{d}_b, \beta, \gamma, \boldsymbol{\omega})$  and  $\boldsymbol{\Psi} = (\tau^2, \kappa^2)$
  - 2 Latent Gaussian field:  $p(\boldsymbol{\alpha}|\boldsymbol{\Psi}) \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_{\boldsymbol{\Psi}})$
  - 3 Hyperparameters:  $|\boldsymbol{\Psi}| = 2$
- We extended INLA implementation (Sauter and Held, 2015) to different NMA models (Jackson model, NMA-regression) and also automation

## Smoking dataset (Hasselblad, 1998)

- 24 trials investigating four interventions to aid smoking cessation
- Coding; 1: no contact, 2: self-help, 3: individual counseling and 4: group counseling
- Area of circle: participants; width of line: trials
- 8 designs, 1-3-4 and 2-3-4 three arm trials

Network Plot



# Specifications for Jackson model

- Basic parameters:  $\mathbf{d}_b = \{d_{12}, d_{13}, d_{14}\}$
- Priors:
  - ⇒ Fixed effects:  $\boldsymbol{\mu}, \mathbf{d}_b \sim \mathcal{N}(0, 1000)$
  - ⇒ Hyperparameters:  $\tau, \kappa \sim \mathcal{U}(0, 5)$ .

## MCMC implementation

- MCMC via JAGS program
- MCMC: 74 parameters to check convergence
- 50,000 iterations after burn-in of 30,000 iterations
- To ensure MCSE below 0.005

## nmaINLA R package

- Publicly available from Github repository

```
library(devtools)
install_github("gunhanb/nmaINLA")
```

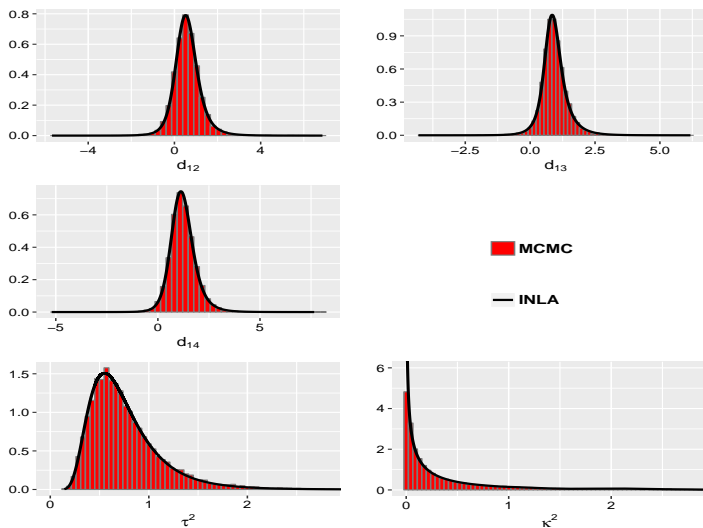
- Data preparation step
- Fitting a Jackson model

```
nma_inla(SmokdatINLA, likelihood = 'binomial', type = 'jackson',
         fixed.par = c(0, 1000), tau.prior = 'uniform', tau.par = c(0, 5),
         kappa.prior = 'uniform', kappa.par = c(0, 5))
```

- MCMC run took 36.3, INLA took 6 seconds.



# Marginal posterior density estimates



# Conclusions

- Common framework for contrast-based NMA models to analyze dataset with different endpoints
- INLA's advantages over MCMC
  - ⇒ Faster, no need to check convergence diagnostics
- `nmaINLA` extracts features needed for NMA
- Arm-based models are also possible, but not implemented yet.

# Outlook

- INLA especially useful when refitting model is needed
  - Extensive simulations
  - Sensitivity analysis (for different priors)
  - Node-splitting method (a NMA technique)
  - Cross-validation (for model selection)

# References I

## Acknowledgements

- Dr. Rafael Sauter

Dias, S., Welton, N. J., Sutton, A. J., and Ades, A. (2011). NICE DSU Technical Support Document 2: A Generalised Linear Modelling Framework for Pairwise and Network Meta-analysis of Randomised Controlled Trials. last updated September 2016.

Hasselblad, V. (1998). Meta-analysis of multitreatment studies. *Medical Decision Making*, 18(1):37–43.

Higgins, J. P. T., Jackson, D., Barrett, J. K., Lu, G., Ades, A. E., and White, I. R. (2012). Consistency and inconsistency in network meta-analysis: concepts and models for multi-arm studies. *Research Synthesis Methods*, 3(2):98–110.

## References II

- Higgins, J. P. T. and Whitehead, A. (1996). BORROWING STRENGTH FROM EXTERNAL TRIALS IN A META-ANALYSIS. *Statistics in Medicine*, 15(24):2733–2749.
- Jackson, D., Barrett, J. K., Rice, S., White, I. R., and Higgins, J. P. (2014). A design-by-treatment interaction model for network meta-analysis with random inconsistency effects. *Statistics in Medicine*, 33(21):3639–3654.
- Lu, G. and Ades, A. E. (2006). Assessing Evidence Inconsistency in Mixed Treatment Comparisons. *Journal of the American Statistical Association*, 101(474):447–459.
- Rue, H., Martino, S., and Chopin, N. (2009). Approximate Bayesian inference for latent Gaussian models by using integrated nested Laplace approximations. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 71(2):319–392.
- Salanti, G. (2012). Indirect and mixed-treatment comparison, network, or multiple-treatments meta-analysis: many names, many benefits, many concerns for the next generation evidence synthesis tool. *Research Synthesis Methods*, 3(2):80–97.

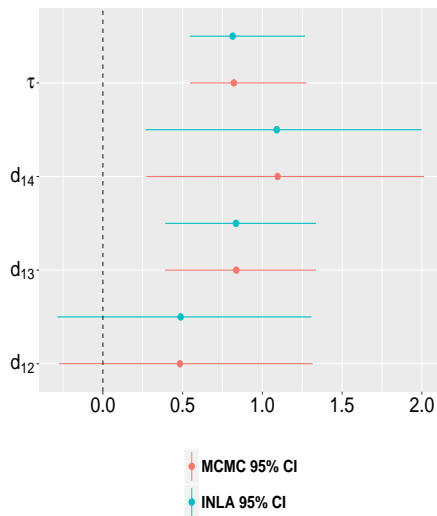
## References III

- Sauter, R. and Held, L. (2015). Network meta-analysis with integrated nested Laplace approximations. *Biometrical Journal*, 57(6):1038–1050.
- van Valkenhoef, G., Tervonen, T., de Brock, B., and Hillege, H. (2012). Algorithmic parameterization of mixed treatment comparisons. *Statistics and Computing*, 22(5):1099–1111.

## Lu-Ades Model (Lu and Ades, 2006)

- Uses cycle-inconsistency approach
- Assumption: inconsistency **only** occurs from 3-cycles
- Basic parameters should form a spanning tree
- Cycle-specific inconsistency random effects:  $\omega_{jkl} \sim \mathcal{N}(0, \kappa^2)$
- Multi-arm trials are inherently consistent
- Number of inconsistency random effects:  $\text{ICDF} = \#\mathbf{d}_f - S$   
where  $S$  is the number of cycles only formed by a multi-arm trial
- Algorithm for ICDF (van Valkenhoef et al., 2012), but not efficient
- In the presence of multi-arm trials, **results depend on treatment ordering!**

# Consistency model (MCMC vs INLA)





# MCMC settings

## Consistency model

- Burn-in: 30.000 iterations
- After burn-in: 20.000 iterations
- 3 chains, 5 thinning parameter
- MCMC run took 29.1, INLA took 2.2 seconds.

## Jackson model

- Burn-in: 30.000 iterations
- After burn-in: 50.000 iterations
- 3 chains, 5 thinning parameter
- To ensure Monte-Carlo standard error is below 0.005 for all parameters

## Approximation error of INLA

- No analytical expression for approximation error of INLA
- But, in (quasi)-complete separation situation (binomial endpoints), INLA shows some inaccuracy
- One way to overcome is by using **weakly informative priors**

# Multi-arm trials

- Convenient and simple variance-covariance matrix of heterogeneity, since we assume
- The homogeneity of between-study variations for every treatment comparison
- Also, for inconsistency, the homogeneity of inconsistency for every treatment comparison

# Inconsistency parameters

| Design | Parameter       | MCMC  |        | INLA  |        |
|--------|-----------------|-------|--------|-------|--------|
|        |                 | Mean  | Stdev. | Mean  | Stdev. |
| 1      | $\omega_{13}^1$ | 0.02  | 0.56   | 0.02  | 0.53   |
|        | $\omega_{14}^1$ | -0.29 | 0.67   | -0.28 | 0.64   |
| 2      | $\omega_{23}^2$ | -0.07 | 0.57   | -0.07 | 0.55   |
|        | $\omega_{24}^2$ | -0.10 | 0.58   | -0.10 | 0.55   |
| 3      | $\omega_{13}^3$ | -0.10 | 0.52   | -0.10 | 0.50   |
| 4      | $\omega_{12}^4$ | -0.13 | 0.58   | -0.13 | 0.55   |
| 5      | $\omega_{14}^5$ | 0.42  | 0.81   | 0.39  | 0.76   |
| 6      | $\omega_{23}^6$ | -0.11 | 0.57   | -0.11 | 0.55   |
| 7      | $\omega_{24}^7$ | 0.09  | 0.57   | 0.09  | 0.55   |
| 8      | $\omega_{34}^8$ | -0.04 | 0.53   | -0.03 | 0.50   |