

A geometric view of overdispersion

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- Models for overdispersed data can be expressed in terms of **simple geometrical operations**

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- We can proceed depending of type/cause of overdispersion by modelling a **more flexible variance function** eg by adding extra parameters

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- **Do not exist** for discrete data! Jorgensen (1997)

Binomial $Y \sim \text{Bin}(n, \mu)$

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Example: Beta-Binomial

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Use simple **dual affine geometry!!** -1 type and -1 type extensions

Overdispersed Binomials: +1 type extensions

$$\begin{aligned} f(y; \mu) &= \exp(y \theta(\mu) - \psi(\theta(\mu))) \nu(y) \\ &\quad \downarrow \\ g(y; \mu, \phi) &= \exp(y \eta + T(y)\phi - \kappa(\eta, \phi)) \nu(y) \end{aligned}$$

where $\mu = E_g[Y]$.

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Family	$T(y)$
Double Binomial	$y \log(y/n) + (n - y) \log((n - y)/n)$
Multiplicative	$-x(n - x)$
Shrink base measure	$-\log \binom{n}{x}$

Overdispersed Binomials: +1 type extensions

- Locally to $\phi = 0$

$$\text{Var}(Y; \mu, \eta) = v(\mu) + \phi \zeta(\mu)$$

where

$$\zeta(\mu) = \text{cov}_f(T(Y), Y^2) - \frac{\text{cov}_f(Y, Y^2)\text{cov}_f(T(Y), Y)}{\text{Var}_f(Y)}$$

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- Parameters η and $\mu = E_g(Y)$ are **orthogonal** (mixed parametrisation)

Double Binomial (Efron, 1986)

Consider the **+1 joining** of $f(y; \mu)$ and $f(y; y)$

$$g(y; \mu, \lambda) = c(\mu, \lambda)[f(y; \mu)]^\lambda [f(y; y)]^{1-\lambda}$$

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Overdispersed Binomials: +1 type extension

Let now $T(y)$ to depend on μ

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Examples

Family	$T(y)$
Overdispersion Test	$\frac{d^2 f(x; \theta(\mu)) / d\mu^2}{f(x; \theta(\mu))}$
Copas & Eguchi (2005)	semiparametric

Overdispersion test (Cox, 1983, Rayner & Best, 2009)

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A **different approach** is to use **local-mixture-like** expressions such as

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Examples

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Local Mixture order 2	$\frac{d^2 f(x; \theta(\mu)) / d\mu^2}{f(x; \theta(\mu))}$
Correlated Binomial	ditto

- **Modified Laplace** expansion

$$\int f(y; m) dH(m) \sim g(y; \mu, \phi) = f(x; \mu) \left[1 + \phi \frac{d^2 f(x; \theta(\mu)) / d\mu^2}{f(x; \theta(\mu))} \right]$$

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- In **general**, let Z_1, \dots, Z_n be dependent Bernoulli(p) random variables such that $\text{Cov}[Z_i, Z_j] = \phi$ for $i \neq j$. If $Y = \sum_{i=1}^n Z_i$

$$\begin{aligned}\text{Var}[Y] &= \sum_{i=1}^n \text{Var}[Z_i] + \sum_{i \neq j} \text{Cov}[Z_i, Z_j] \\ &= v(\mu)[1 + (n - 1)\rho]\end{aligned}$$

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The sum $Y = \sum_{i=1}^n Z_i$ has density

$$g(y; \mu, \phi) = \frac{\alpha^y (1 - \alpha)^{m-1-2y} \beta^{y+1}}{\alpha + \beta} S(y; \alpha, \beta, n)$$

$$\begin{aligned} E[Y] &= \mu \\ \text{Var}[Y] &= v(\mu)[1 + (n-1)r(\alpha, \beta)] \end{aligned}$$

Overdispersed Binomial: Mixed parametrisation

Embed the extended binomial $g(y; \mu, \phi)$ into the **multinomial** via

$$g(y; \mu, \phi) \mapsto (f(0; \mu, \phi), f(1; \mu, \phi), \dots, f(n; \mu, \phi))^T$$

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so density of $\mathbf{U} = (U_1, \dots, U_N)$ where $U_i = \mathbb{I}(Y = i)$ is

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where $\boldsymbol{\chi}(\mu, \phi) =$

$(\log(f(1, \mu, \phi)/f(0, \mu, \phi)), \dots, \log(f(n, \mu, \phi)/f(0, \mu, \phi)))$. Define

$$\mathbf{d} = \begin{pmatrix} 1 \\ 2 \\ \vdots \\ N \end{pmatrix} \quad \mathbf{d}_T = \begin{pmatrix} T(1) \\ T(2) \\ \vdots \\ T(N) \end{pmatrix} \quad \mathbf{c} = \begin{bmatrix} \log \binom{N}{1} \\ \log \binom{N}{2} \\ \vdots \\ \log \binom{N}{N} \end{bmatrix}$$

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Now use **mixed parametrisation**

$$(\mu, \mu^{(2)}, \boldsymbol{\omega}^*(\mu, \phi)) \quad \text{or} \quad (\mu, \text{Var}(Y), \boldsymbol{\omega}^*) \quad \boldsymbol{\omega}^* \in \mathbb{R}^{n-2}$$

to **characterise** the extended Binomial $g(y; \mu, \phi)$

Overdispersed Binomial: Mixed parametrisation

Family	$Var(Y)$	ω^*
Binomial	$v(\mu)$	ω_0^*
Beta-Binomial	$v(\mu)\phi$	$\omega_0^* + \mathbf{u}(\mu, \phi)$
Multiplicative	$v(\mu)\phi$	ω_0^*
Multiplicative+	$Var(Y)$	$\omega_0^* + \delta\mathbf{u}$
Markov	$v(\mu)\phi$	$\omega_0^* + \mathbf{v}(\mu, \phi)$

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Summary

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- Goes **beyond exponential and mixture** family structures

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