

Comparisons theorems for slice sampling

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Algorithms seminar, Warwick

30th January 2024



- 1 Introduction: MCMC
 - MCMC
 - Slice Sampling
- 2 Convergence of Markov chains and comparisons
- 3 Comparisons of Slice Sampling
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Seek **learn** or **infer** values of the **parameter** x which are **commensurate** with the observed dataset y .

The Bayesian approach

Encode prior beliefs into a **prior distribution** $\nu(x)$, and define **likelihood** $\ell_y(x) := f_x(y)$.

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We are then interested in quantities of the form

$$I = \pi(f) = \int_{\mathcal{X}} f(x)\pi(x) dx,$$

e.g. $f(x) = \|x\|^p$ (**posterior moments**), $f(x) = 1_A(x)$ (**credible sets / posterior tail probabilities**), etc.

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So instead, approximate I by **sampling** $X_1, X_2, \dots, X_n \sim \pi$ and consider

$$I_n := \frac{1}{n} \sum_{i=1}^n f(X_i) \approx I = \int_{\mathcal{X}} f(x)\pi(x) dx.$$

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We simulate a π -reversible ergodic Markov chain,

$$X_1, X_2, \dots$$

where $X_n \rightarrow \pi$ in distribution and considering

$$I_n := \frac{1}{n} \sum_{i=1}^n f(X_i) \approx I = \int_{\mathcal{X}} f(x) \pi(x) dx.$$

Algorithm 1 Metropolis–Hastings (MH)

```
1: initialise:  $X_0 = x_0, i = 0$ 
2: while  $i < N$  do
3:    $i \leftarrow i + 1$ 
4:   simulate  $Y_i \sim Q(X_{i-1}, \cdot)$ 
5:    $\alpha(X_{i-1}, Y_i) = 1 \wedge \frac{q(Y_i, X_{i-1})\pi(Y_i)}{q(X_{i-1}, Y_i)\pi(X_{i-1})}$ 
6:   with probability  $\alpha(X_{i-1}, Y_i)$ 
7:      $X_i \leftarrow Y_i$ 
8:   else
9:      $X_i \leftarrow X_{i-1}$ 
10: return  $(X_i)_{i=1, \dots, N}$ 
```

Random walk Metropolis

Popular approach: [Random Walk Metropolis](#) (RWM) [[Metropolis et. al. \(1953\)](#)]:

$$Q(X_{i-1}, \cdot) = \mathcal{N}(X_{i-1}, \sigma^2 \cdot \mathbf{I}).$$

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[Slice Sampling](#) was introduced to try and circumvent these deficiencies.

Slice Sampling [Neal (2003)]

Target: $\pi(\mathrm{d}x) = \varpi(x)\nu(\mathrm{d}x)$ on \mathcal{X} , reference measure ν .

$$G(t) := \{x \in \mathcal{X} : \varpi(x) > t\}, \quad \nu_t := \frac{\nu(\cdot \cap G(t))}{\nu(G(t))}.$$

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Algorithm 3 Ideal Slice Sampling

```
1: initialise:  $X_0 = x_0, i = 0$ 
2: while  $i < N$  do
3:    $i \leftarrow i + 1$ 
4:   Sample  $t \sim \text{Unif}([0, \varpi(x)])$ ;
5:   Sample  $Y \sim \nu_t$ ;
6:   Set  $X_i = Y$ .
7: return  $(X_i)_{i=1, \dots, N}$ 
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But this can be relaxed, if instead we have access to ν_t -reversible kernels!

(Similar to going from Gibbs sampling \rightsquigarrow Metropolis-within-Gibbs.)

Hybrid Slice Sampling

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Algorithm 5 Hybrid Slice Sampling (HSS)

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2: while  $i < N$  do
3:    $i \leftarrow i + 1$ 
4:   Sample  $t \sim \text{Unif}([0, \varpi(x)])$ ;
5:   Sample  $Y \sim H_t(X_{i-1}, \cdot)$ ;
6:   Set  $X_i = Y$ .
7: return  $(X_i)_{i=1, \dots, n}$ 
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This is still a π -reversible Markov chain.

Hybrid Slice Sampling examples

HSS also defines a π -reversible Markov chain.

- Random Walk Metropolis on the slice
- Hit-and-Run
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Question: how much worse?

Example: Metropolis chains

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For example: $\nu = \text{Leb}$, $Q(x, \cdot) = \mathcal{N}(x, \sigma^2 \cdot \mathbf{I})$, **Random Walk Metropolis**.
 $\nu = \mathcal{N}(0, \mathbf{C})$, $Q(x, \cdot) = \mathcal{N}(\rho x, (1 - \rho^2) \cdot \mathbf{C})$, **preconditioned Crank–Nicolson**.

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It turns out that such chains can actually be seen as examples of **Hybrid Slice Samplers**.

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Thus our subsequent results allow us to study such chains.

Summary

We will present a [comparison result](#) which can quantify precisely how much worse the HSS is compared to the ISS.

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We will present a **comparison result** which can quantify precisely how much worse the HSS is compared to the ISS.

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We will do this comparison using the framework of **weak Poincaré inequalities**.

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Intuitively, it relates the convergence of HSS to the convergence rate of ISS in terms of the **convergence rates of the (H_t) kernels**.

We will do this comparison using the framework of **weak Poincaré inequalities**.

This will enable us to give **quantitative bounds**, covering cases when there is **no spectral gap**, i.e. the convergence is **subgeometric**, substantially extending previous results [Qin et. al. (2023), Łatuszyński & Rudolf (2014)].

Overview

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$$\|P^n f - \pi(f)\|_2 \leq (1 - \gamma)^n \|f\|_2, \quad \forall f \in L^2(\pi).$$

However some chains have 0 spectral gap and have only subgeometric convergence;

$$\|P^n f - \pi(f)\|_2^2 \leq \gamma(n) \Phi(f),$$

where

$$\Phi(f) = \|f\|_{\text{osc}}^2 = (\text{ess sup } f - \text{ess inf } f)^2.$$

Standard Poincaré inequalities

We work on $L^2(\mu) = \{f : \mathcal{X} \rightarrow \mathbb{R} : \|f\|_2^2 < \infty\}$, $\langle f, g \rangle := \int fg \, d\mu$,
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For a μ -invariant Markov transition kernel P with $L^2(\mu)$ -adjoint P^* , consider the Dirichlet form $\mathcal{E}(P^*P, f)$, for $f \in L_0^2(\mu)$:

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Standard Poincaré inequality (SPI)

A SPI holds if there exists a constant $C_P > 0$ such that for all $f \in L_0^2(\mu)$,

$$C_P \|f\|_2^2 \leq \mathcal{E}(P^*P, f).$$

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Theorem (Geometric convergence)

Under a *standard Poincaré inequality*, we have for all $f \in L_0^2(\mu)$, $n \in \mathbb{N}_0$,

$$\|P^n f\|_2^2 \leq (1 - C_P)^n \|f\|_2^2.$$

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Proof. Rewriting the SPI, see $\mathcal{E}(P^*P, f)$ behaves like a discrete derivative:

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$$\begin{aligned} C_P \|f\|_2^2 &\leq \mathcal{E}(P^*P, f) = \|f\|_2^2 - \langle P^*P f, f \rangle \\ &= \|f\|_2^2 - \|P f\|_2^2 \\ \Rightarrow \|P f\|_2^2 &\leq (1 - C_P) \|f\|_2^2. \end{aligned}$$

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The rest is by induction. \square

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A **WPI** holds if: for some such β , Φ , $\forall s > 0$, $f \in L_0^2(\mu)$,

$$\|f\|_2^2 \leq s \mathcal{E}(P^*P, f) + \beta(s) \Phi(f).$$

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E.g. $\beta(s) = c_0 s^{-c_1}$.

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Define

$$K(u) := u\beta(1/u), \quad u \geq 0,$$

$$K^*(v) := \sup_{u \geq 0} \{uv - K(u)\}, \quad v \geq 0,$$

$$F(x) := \int_x^1 \frac{dv}{K^*(v)}, \quad 0 < x \leq 1.$$

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Theorem ([Andrieu, Lee, Power, W. (2022)])

Under a *weak Poincaré inequality*, we have, $\forall n \in \mathbb{N}_0, f \in L_0^2(\mu)$,

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Examples

Theorem ([Andrieu, Lee, Power, W. (2022)])

Under a *weak Poincaré inequality*, we have, $\forall n \in \mathbb{N}_0, f \in L_0^2(\mu)$,

$$\|P^n f\|_2^2 \leq \Phi(f) F^{-1}(n).$$

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Intuition: the *faster* β decays, the *faster* the rate of convergence.

Comparisons of Markov chains

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In [Andrieu, Lee, Power, W. (2022)], we used this machinery to study pseudo-marginal MCMC.

Comparisons of Markov chains II

Suppose we are able to establish the comparison: $\forall s > 0, f \in L_0^2(\mu)$,

$$\mathcal{E}(\textcolor{blue}{P}, f) \leq s \mathcal{E}(\textcolor{red}{\tilde{P}}, f) + \textcolor{red}{\beta}(\textcolor{red}{s}) \Phi(f).$$

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from which a convergence bound for \tilde{P} can be obtained.

- 1 Introduction: MCMC
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Main result

We have $\pi(dx) = \varpi(x)\nu(dx)$

$$m_t := \nu(G(t)),$$

recall $\nu_t = \nu(\cdot \cap G(t))/m_t$. Let's write U for Ideal Slice Sampling, H for Hybrid Slice Sampling.

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$$\mathcal{E}(U, f) \leq s \cdot \mathcal{E}(H, f) + \beta(s)\Phi(f),$$

where $\beta : (0, \infty) \rightarrow [0, \infty)$ is given by

$$\beta(s) := \int_0^{\|\varpi\|_\infty} \beta_t(s) m_t dt,$$

$$\mathcal{E}(U, f) \leq s \cdot \mathcal{E}(H, f) + \beta(s) \Phi(f), \quad \beta(s) := \int_0^{\|\pi\|_\infty} \beta_t(s) m_t dt.$$

We see that the convergence rate of U is a kind of weighted average of the convergence rates of each H_t .

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This significantly extends the previous work of [Łatuszyński & Rudolf (2014)].

Example: Independent Metropolis–Hastings (IMH)

Recall the $\text{IMH}(\pi, q)$ chain: at each iteration with $X_n = x$, propose $Y \sim q$, and accept this proposal with probability

$$\alpha(x, Y) = 1 \wedge \frac{\pi(Y)q(x)}{\pi(x)q(Y)}.$$

It is known that the $\text{IMH}(\pi, q)$ satisfies an SPI with constant $\|\mathrm{d}\pi/\mathrm{d}q\|_\infty^{-1}$ [Mengersen and Tweedie (1996)].

Explicit example: Exponential distributions

Consider the case when $\mathcal{X} = [0, \infty)$, $\pi(x) = \exp(-x)$, and $\nu(dx) = \lambda \exp(-\lambda x) dx$ for some $\lambda > 0$.

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Lemma

This Ideal Slice Sampler has a spectral gap which can be lower-bounded explicitly:

- When $\lambda \in (0, 1)$, the slice sampler has a spectral gap of at least $\frac{1+\lambda}{2}$;
- When $\lambda > 1$, the slice sampler has a spectral gap of at least $(2\lambda - 1)^{-2}$.

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Derived using a [contractivity](#) argument.

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Consider doing now performing an IMH with proposal ν on each slice:

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When $\lambda \in (0, 1)$, we obtain the following comparison for the IMH:

$$\beta(s) = \frac{1}{4} + \frac{1}{4}(1 - s^{-1})^{1/\lambda} \left(\frac{1}{\lambda(s-1)} - 1 \right) \sim s^{-1},$$

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When $\lambda > 1$, we obtain the following comparison for the IMH:

$$\beta(s) = \frac{\lambda - 1}{4\lambda} s^{-1/\lambda},$$

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Thus we arrive at the comparison

$$\mathcal{E}(U, f) \leq (2^{33} \kappa^2 d^2) \mathcal{E}(H, f).$$

I.e. **spectral gap of H** is at least $2^{-33} \kappa^{-2} d^{-2}$ times the spectral gap of U .

Overview

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Preprint forthcoming!

Thanks for listening! I



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Thanks for listening! II



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Convergence of MCMC

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Theorem ([?, ?])

*RWM converges to equilibrium **exponentially** fast if* and only if π has an **exponential moment** (e.g. $\pi(x) \propto \exp(-\|x - \mu\|^\alpha)$, $\alpha \geq 1$). Otherwise, the chain converges at a **subgeometric** (e.g. **polynomial**) rate.*

L^2 convergence and Dirichlet forms

We work on $L^2(\pi) = \{f : \mathcal{X} \rightarrow \mathbb{R} : \|f\|_2^2 < \infty\}$, $\langle f, g \rangle := \int fg \, d\pi$,
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For a π -invariant Markov transition kernel P with $L^2(\pi)$ -adjoint P^* , define the Dirichlet form $\mathcal{E}(P^*P, f)$, for $f \in L_0^2(\pi)$:

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So it will be sufficient to lower bound $\mathcal{E}(P, f)$.

Lemma ([?])

For nonconstant nonnegative $g \in L_0^2(\pi)$, we have the lower bound

$$\mathcal{E}(P, g) \geq \text{Var}_\pi(g) \cdot \frac{1}{2} \cdot \Lambda_P \left(\frac{4[\pi(g)]^2}{\text{Var}_\pi(g)} \right),$$

where Λ_P is the spectral profile of P .

Lemma

For π -reversible P , we have the further lower bound

$$\Lambda_P(v) \geq \begin{cases} \frac{1}{2} \Phi_P(v)^2 & 0 < v \leq 1/2, \\ \frac{1}{2} [\Phi_P^*]^2 & v > 1/2. \end{cases}$$

Proof of convergence bound (I)

Fix $f \in L_0^2(\mu)$. Have that

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Set $u := 1/s$, $K(u) := u\beta(1/u)$.

$$\frac{\mathcal{E}(P^*P, f)}{\Phi(f)} \geq u \cdot \frac{\|f\|_2^2}{\Phi(f)} - K(u), \quad \forall u > 0.$$

$$\frac{\mathcal{E}(P^*P, f)}{\Phi(f)} \geq \sup_{u>0} \left\{ u \cdot \frac{\|f\|_2^2}{\Phi(f)} - K(u) \right\} =: K^* \left(\frac{\|f\|_2^2}{\Phi(f)} \right).$$

Call this final inequality optimized WPI (oWPI).

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Now define

$$F(x) := \int_x^1 \frac{dv}{K^*(v)}, \quad x \in (0, a], \quad h_n := \frac{\|P^n f\|_2^2}{\Phi(f)}.$$

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So we invert this to obtain

$$\|P^n f\|_2^2 \leq \Phi(f) F^{-1}(n). \quad \square$$