

# Monte Carlo Filtering of Piecewise-Deterministic Processes

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EPSRC Workshop on MCMC and Related Methods

Background

Piecewise  
Deterministic  
Processes

Sequential Monte  
Carlo (Samplers)

Methodology

Particle Filtering of  
PDPs

Auxiliary Particle  
Filtering of PDPs

Examples

Object Tracking

Rate Estimation: Shot  
Noise Cox Processes

Summary

# Outline

- ▶ Background
  - ▶ Piecewise Deterministic Processes (PDPs)
  - ▶ Sequential Monte Carlo (Samplers)
- ▶ Methodology
  - ▶ “Particle Filtering” of PDPs
  - ▶ “Auxiliary Particle Filtering” of PDPs
- ▶ Examples
  - ▶ Object Tracking
  - ▶ Rate Estimation: Shot Noise Cox Processes

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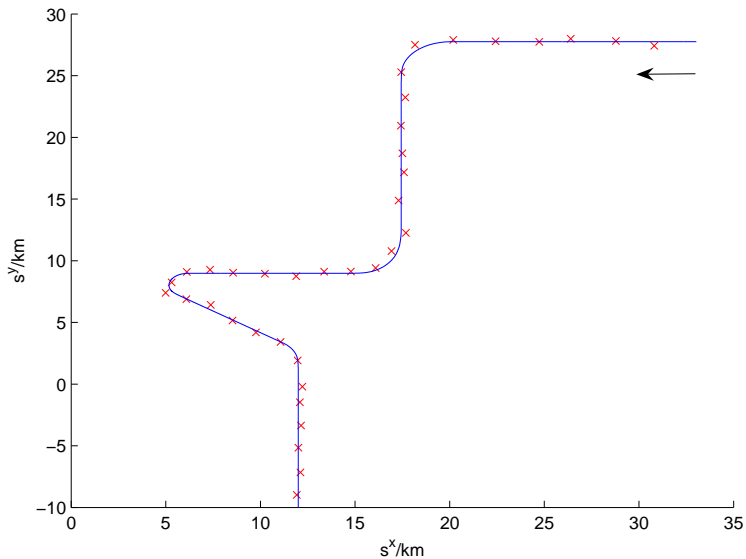
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# Motivation:

## Observing a Manoeuvring Object

- ▶ For  $t \in \mathbb{R}_0^+$ , consider object with position  $s_t$ , velocity  $v_t$  and acceleration  $a_t$
- ▶ Summarise state by  $\zeta_t = (s_t, v_t, a_t)$
- ▶ From initial condition  $\zeta_0$ , state evolves until random time  $\tau_1$ , at which acceleration jumps to a new random value, yielding  $\zeta_{\tau_1}$
- ▶ From  $\zeta_{\tau_1}$ , evolution until  $\tau_2$ , state becomes  $\zeta_{\tau_2}$ , etc.
- ▶ Observation times,  $(t_n)_{n \in \mathbb{N}}$ , at each  $t_n$  a noisy measurement of the object's position is made

# Example Trajectory



Monte Carlo  
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# An Abstract Formulation

- ▶ Pair Markov chain  $(\tau_j, \theta_j)_{j \in \mathbb{N}}$ ,  $\tau_j \in \mathbb{R}^+$ ,  $\theta_j \in \Theta$

$$p(d(\tau_j, \theta_j) | \tau_{j-1}, \theta_{j-1}) = q(d\theta_j | \theta_{j-1}, \tau_j, \tau_{j-1}) f(d\tau_j | \tau_{j-1}),$$

- ▶ Count the jumps  $\nu_t := \sum_j \mathbb{I}_{[\tau_j \leq t]}$
- ▶ Deterministic evolution function  $F : \mathbb{R}_0^+ \times \Theta \rightarrow \Theta$ , s.t.  
 $\forall \theta \in \Theta$ ,

$$F(0, \theta) = \theta$$

- ▶ Signal process  $(\zeta_t)_{t \in \mathbb{R}_0^+}$ ,

$$\zeta_t := F(t - \tau_{\nu_t}, \theta_{\nu_t})$$

- ▶ This describes a piecewise deterministic process.
- ▶ It's partially observed via observations  $(Y_n)_{n \in \mathbb{N}}$ , with likelihood function  $g_n(y_n | \zeta_{t_n})$

# Sequential Importance Resampling

For a sequence,  $\pi_n$ , of distributions over spaces  $E^n$ :

- ▶ Set  $n = 1$
- ▶ For  $i = 1 : N$ 
  - ▶ Sample  $X_1^i \sim q_1$ .
  - ▶ Calculate  $W_1^i \propto \pi_1(X_1^i)/q_1(X_1^i)$ .
- ▶ Iterate:  $n \leftarrow n + 1$ 
  - ▶ Resample  $(W_{n-1}^i, X_{n-1}^i)$  to obtain  $(1/N, X_{n,1:n-1}^i)$ .
  - ▶ For  $i = 1 : N$ 
    - ▶ Sample  $X_{n,n}^i \sim q_n(\cdot | X_{n,1:n-1}^i)$ .
    - ▶ Calculate

$$W_n^i \propto \frac{\pi_n(X_{n,1:n}^i)}{\pi_{n-1}^i(X_{n,1:n-1}^i)q_n(X_{n,n}^i | X_{n,1:n-1}^i)}$$

In a filtering context, can use  $\pi_n(x_{1:n}) = p(x_{1:n} | y_{1:n})$ .

# The Variable Rate Particle Filter

The VRPF iteration:

- ▶ Resample  $(W_{n-1}^i, (k_{n-1}^i, \tau_{1:k_{n-1}^i}^i, \theta_{1:k_{n-1}^i}^i))$ .
- ▶ For  $i = 1 : N$ 
  - ▶ While  $\tau_{n,k_n}^i < t_n$ 
    - ▶ Sample  $\tau_{n,k_n+1}^i \sim f(\cdot | \tau_{n,k_n}^i)$ .
    - ▶  $k_n^i \leftarrow k_n^i + 1$ .
    - ▶ Sample  $\theta_{n,k_n^i}^i \sim h(\cdot | \tau_n, \theta_{n,k_{n-1}^i}^i)$ .
  - ▶ Calculate

$$W_n^i \propto \frac{g(y_n | x_n^i) f(\theta_{k_{n-1}^i+1:k_n^i} | \theta_{k_{n-1}^i})}{h(\theta_{k_{n-1}^i+1:k_n^i} | \tau_{n,1:k_{n-1}^i}^i, \theta_{n,k_{n-1}^i}^i)}$$

Something like the bootstrap filter.

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# SMC Samplers

SMC can be used to sample from *any* sequence of distributions (Del Moral et al., 2006).

- ▶ Given *target* distributions,  $\eta_n$ , on  $E_n \dots$ ,
- ▶ construct a synthetic sequence  $\tilde{\eta}_n$  on spaces  $\bigotimes_{p=1}^n E_p$
- ▶ by introducing Markov kernels,  $L_p$  from  $E_{p+1}$  to  $E_p$ :

$$\tilde{\eta}_n(x_{1:n}) = \eta_n(x_n) \prod_{p=1}^{n-1} L_p(x_{p+1}, x_p),$$

- ▶ These distributions
  - ▶ have the target distributions as time marginals,
  - ▶ have the correct structure to employ SMC techniques.



# Inference and PDPs

- ▶ Consider the spaces  $(E_n)_{n \in \mathbb{N}}$ ,

$$E_n = \bigsqcup_{k=0}^{\infty} \{k\} \times \mathbb{T}_{n,k} \times \Theta^{k+1}$$

$$\mathbb{T}_{n,k} = \{\tau_{1:k} : 0 < \tau_1 < \tau_2 < \dots < \tau_k \leq t_n\}.$$

- ▶ Define  $k_n := \nu_{t_n}$  and  $X_n = (\zeta_0, k_n, \tau_{1:k_n}, \theta_{1:k_n}) \in E_n$
- ▶ Sequence of posterior distributions  $(\eta_n)_{n \in \mathbb{N}}$

$$\begin{aligned} \eta_n(x_n) &\propto q(\zeta_0) \prod_{j=1}^{k_n} f(\tau_j | \tau_{j-1}) q(\theta_j | \theta_{j-1}, \tau_j, \tau_{j-1}) \\ &\quad \times \prod_{p=1}^n g_p(y_p | \zeta_{t_p}) S(\tau_{k_n}, t_n) \end{aligned}$$

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# Particle Filtering of PDPs

- ▶ Can use SMC sampler to target  $\eta_n$  for  $n = 1, \dots$
- ▶ This allows more sophisticated proposals:
  - ▶ Use of observations.
  - ▶ Explicit dimension-changing proposals.
  - ▶ No sampling into the future.
- ▶ Often interested only in  $p(\zeta_{t_n} | y_{1:n})$ .
- ▶ Consequently, need only keep track of  $(k_n, \tau_{k_n}, \theta_{k_n})$ .
- ▶ Many techniques from particle filtering can be used.

# Algorithmic Details

- ▶ Use a mixture of moves:
  - ▶ *Birth* moves.
  - ▶ *Refinement* moves.
- ▶ Augment with MCMC moves if required:
  - ▶ Can include RJMCMC moves.
- ▶ Use observations to obtain good proposals.
  
- ▶ Recover the VRPF under certain circumstances.

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# Auxiliary Particle Filtering of PDPs

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- ▶ Can construct an analogue of discrete time APF:
  - ▶ Use  $n + 1^{\text{st}}$  observation at time  $n$ .
  - ▶ Prevents resampling eliminating promising particles.
  - ▶ Pre-weighting is then corrected for by later weights.
  - ▶ Interpretation:
    - ▶ Approximate  $\hat{p}(\cdot|y_{1:n+1})$  at time  $n$ .
    - ▶ Correct with weights  $p(\cdot|y_{1:n})/\hat{p}(\cdot|y_{1:n+1})$ .

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- ▶ An auxiliary PDP particle filter comprises:
  - ▶ SMC sampler for auxiliary sequence  $(\mu_n)_{n \geq 1}$ ,
  - ▶ Sequence of importance weights,  $(\widetilde{W}_n)$  between  $(\mu_n)_{n \geq 1}$  and  $(\eta_n)_{n \geq 1}$ .
- ▶ We use auxiliary distributions of the following form:

$$\mu_n(x_n) \propto V_n(\tau_{n,k_n}, \theta_{n,k_n}, y_{n+1}) \eta_n(x_n)$$

with  $V_n : \mathbb{R}^+ \times \Xi \rightarrow (0, \infty)$ .

- ▶  $V_n$  approximates predictive likelihood.

# Back to the motivating example...

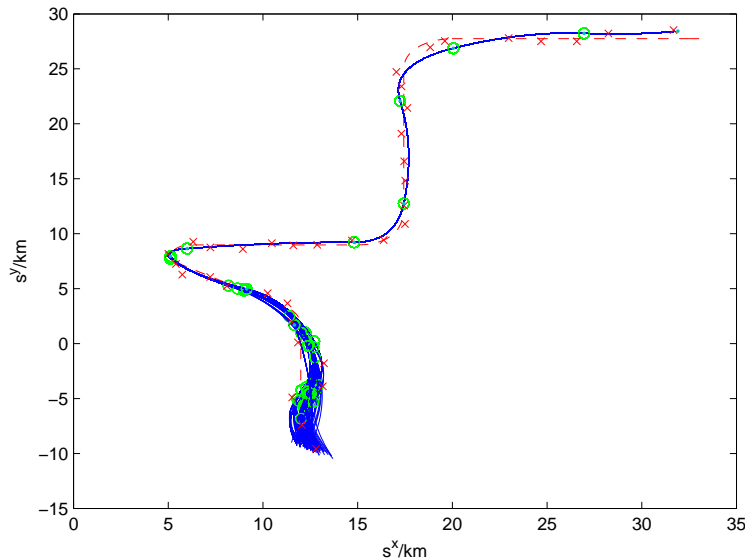
- ▶ Observation interval,  $\Delta = 5$  s.
- ▶ Changepoint intervals are *a priori* Gamma( $a, b$ ) with:

$$a = 10$$

$$b = 5\Delta/a$$

- ▶ Used a mixture of birth and adjustment moves (in proportion 2:1).
- ▶ Observations are noisy range/bearing pairs.
- ▶ Model linearisation used to approximate optimal kernels.
- ▶ Stratified resampling when  $ESS < N/2$ .
  
- ▶ Can also consider a Rao-Blackwellised version.

# The final sample set



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$N$	Godsill et al. 2007		Whiteley et al. 2007	
	RMSE / km	CPU / s	RMSE / km	CPU / s
50	42.62	0.24	0.88	1.32
100	33.49	0.49	0.66	2.62
250	22.89	1.23	0.54	6.56
500	17.26	2.42	0.51	12.98
1000	12.68	5.00	0.50	26.07
2500	6.18	13.20	0.49	67.32
5000	3.52	28.79	0.48	142.84

Root mean square filtering error and  
CPU time — over 200 runs.

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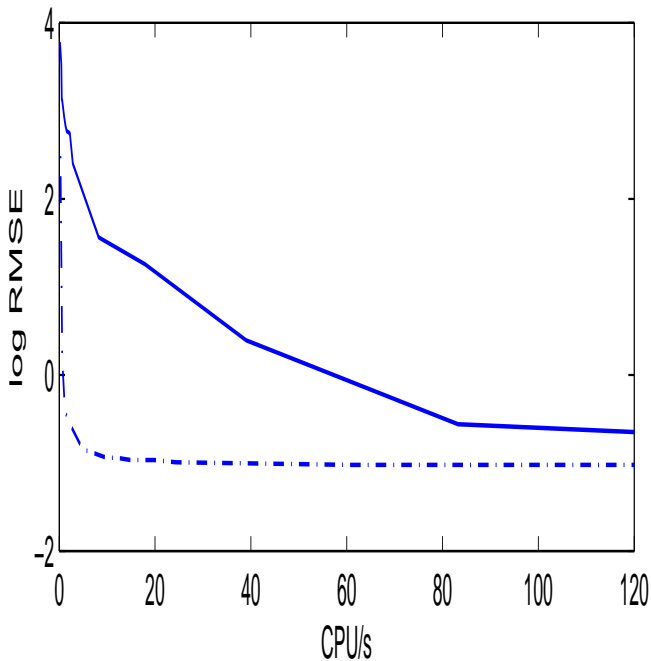
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# Estimation of Insurance Event Rates

The model is specified by the following distributions:

$$f(\tau_j | \tau_{j-1}) = \lambda_\tau \exp(-\lambda_\tau(\tau_j - \tau_{j-1})) \times \mathbb{I}_{[\tau_{j-1}, \infty)}(\tau_j),$$

$$q(\zeta_0) = \exp(-\lambda_\theta \zeta_0) \times \mathbb{I}_{[0, \infty)}(\zeta_0),$$

$$q(\theta_j | \theta_{j-1}, \tau_j, \tau_{j-1}) = \lambda_\theta \exp(-\lambda_\theta(\theta_j - \zeta_{\tau_j}^-)) \times \mathbb{I}_{[\zeta_{\tau_j}^-, \infty)}(\theta_j),$$

where

$$\zeta_{\tau_j}^- = \theta_{j-1} \exp(-\kappa(\tau_j - \tau_{j-1})), \text{ and}$$

$$F(t - \tau, \theta) = \theta \exp(-\kappa(t - \tau)).$$

The likelihood function is given by:

$$g(y_n | \zeta_{(t_{n-1}, t_n]}) = \exp\left(-\int_{t_{n-1}}^{t_n} \zeta_s ds\right) \prod_i \zeta_{y_n, i}$$

where  $y_{n, i}$  is the time of the  $i$ th event observed in  $(t_{n-1}, t_n]$ .

# Implementation Details

- ▶ Proposal step: simple birth proposal.

$$\begin{aligned} K_n(x_{n-1}, dx_n) &= \delta_{k_{n-1}+1}(k_n) \delta_{\tau_{n-1}, 1:k_{n-1}}(d\tau_{n, 1:k_n-1}) \\ &\quad \times \delta_{\theta_{n-1}, 1:k_{n-1}}(d\theta_{n, 1:k_n-1}) h_n(d\tau_{n, k_n}) \\ &\quad \times \pi_n(d\theta_{n, k_n} | x_n \setminus \theta_{n, k_n}), \end{aligned}$$

- ▶ Systematic resampling applied when  $ESS < 0.4N$ .
- ▶ RJMCMC move applied.

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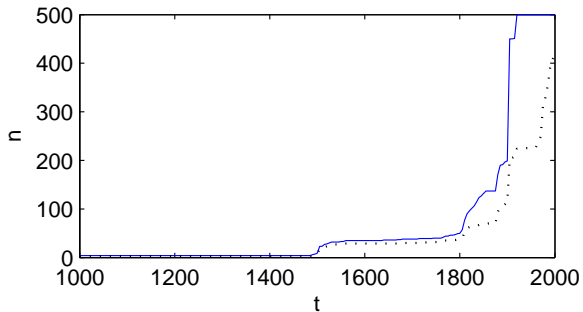
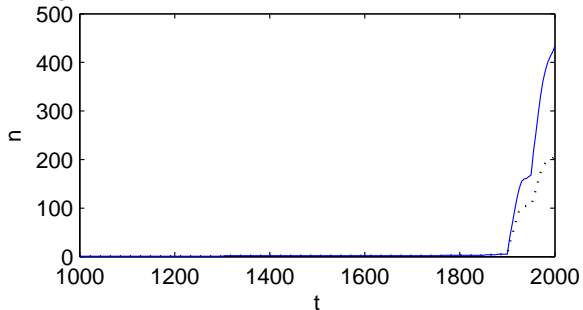
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# Unique Particles



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# SMCTC: A C++ Template Class

- ▶ Implementing SMC algorithms in C/C++ isn't hard.
- ▶ Software for implementing general SMC algorithms.
- ▶ C++ element largely confined to the library.
- ▶ Available (under a GPL-3 license from)  
`www2.warwick.ac.uk/fac/sci/statistics/staff/  
academic/johansen/smctc/`  
or type "smctc" into google.
- ▶ Particle filters can also be implemented easily.

# Conclusions

- ▶ SMC sampler-type approaches allow efficient online algorithms to be developed.
  - ▶ The proposed approach can dramatically outperform existing algorithms.
  - ▶ This approach allows a class of piecewise deterministic models to be used when online inference is required.
  - ▶ The VRPF can be interpreted within this framework, allowing existing convergence results to be applied to it.
- 
- ▶ Possible extensions:
    - ▶ Parameter estimation via PMCMC.

# References



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