

# Let's chop them up!

## (A brief survey on SIR techniques)

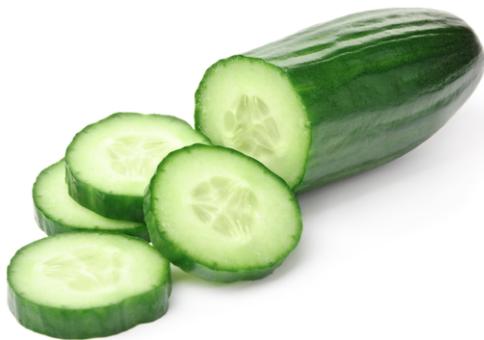
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# This ain't a culinary lecture!



# Outline

- 1 Sliced inverse regression
- 2 The Bayesian partition model
- 3 Other recent developments [optional]
- 4 Concluding remarks

# Background

## The challenge

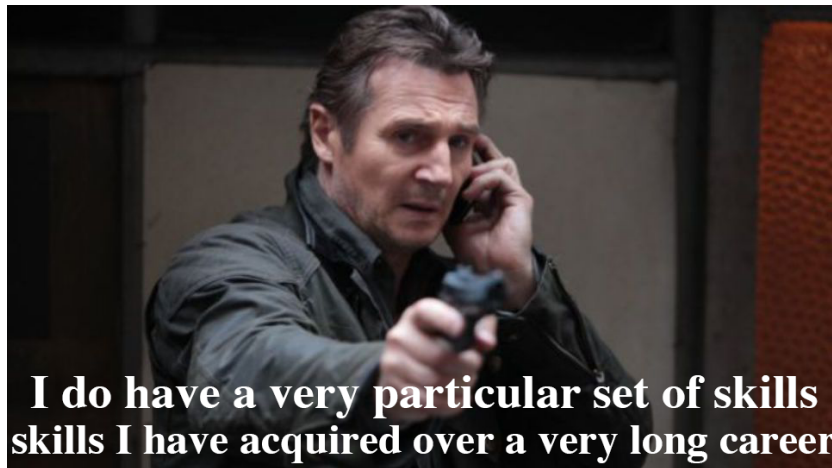
Say  $X$  are some high-dimensional predictors and  $Y$  are some responses of interest, one would like to have a low-dimensional summary  $\tilde{X}$  of  $X$  that is informative about  $Y$ .

## Examples

- $X$ : genetic makeup,  $Y$ : disease risk
- $X$ : historic quotes on stocks,  $Y$ : future prices
- $X$ : brain activations,  $Y$ : psychological status

## Potential gain

- Better model generalizability and interpretability
- More efficient computations



**I do have a very particular set of skills  
skills I have acquired over a very long career**

# Common solutions

## Two summarizing strategies

- Dimension reduction:  $\tilde{X}$  is a transformation of  $X$ 
  - CCA, PLS, RRR
- Variable selection:  $\tilde{X}$  is a subset of  $X$ 
  - LASSO, penGAM, ISIS (not those terrorists)

## Measuring informativeness

- Parametric measures
  - Predictive power of  $\tilde{X}$  on  $Y$
  - Model consistency (likelihood)
  - Association
- Nonparametric measure
  - Shared information

# Aren't they good enough?

## Limitations

- Validity of the model assumptions
- Data consuming
- Computationally challenging
  - Applies to both para. and non-para. solutions

Any more appealing alternatives?



# Regression revisited

## Forward regression

$$\mathbb{E}[Y|X] = \phi(X)$$

- Estimate  $\hat{\phi}_n$  with empirical sample  $(\mathbf{X}_n, \mathbf{Y}_n)$

## Cons

- The family of  $\phi$  may not be known *a priori*
- Estimation often relies on the distribution of  $Y = \psi(X, E)$ 
  - $\psi$  the data generating mechanism
  - $E$  the randomness involved

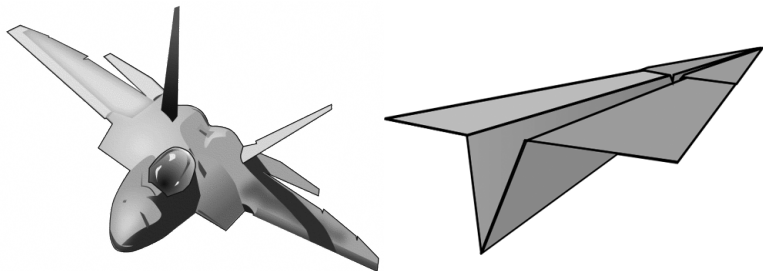
## The catch

- We don't really need  $\phi$  to characterize the dependency
- And we do not need to know the distribution of  $Y$  either



## An simple analogy

- You learn the basic laws of aerodynamics from a paper plane
- But it takes a lot more to build an F22 raptor
- Basics is suffice for us, let's stick with it!!!



# Sliced inverse regression (SIR)

## Inverse regression

$$\mathbb{E}[X|Y] = \eta(Y)$$

Assuming the following general data generation mechanism

$$Y = \psi(X^\top \beta_1, \dots, X^\top \beta_K, E). \quad (1)$$

## Theorem (Li, 1991)

*Under model (1), and assume  $X$  follows elliptical distributions, the centered inverse regression curve  $\bar{\eta}(Y) = \mathbb{E}[X|Y] - \mathbb{E}[X]$  is contained in the linear subspace spanned by  $\Sigma_{XX}\beta_k$  ( $k = 1, \dots, K$ ), where  $\Sigma_{XX}$  denotes the covariance matrix of  $X$ .*

## Sketch of proof.

$$\begin{aligned}
 \mathbb{E}[X|Y] &= \mathbb{E}[\mathbb{E}[X|\eta^T X, Y]|Y] \\
 &= \mathbb{E}[\mathbb{E}[X|\eta^T X]|Y] \\
 &= \mathbb{E}[\mathbb{E}[P_\eta X + Q_\eta X|\eta^T X]|Y] \\
 &= \mathbb{E}[P_\eta X|Y] + \mathbb{E}[\mathbb{E}[Q_\eta X|\eta^T X]|Y]
 \end{aligned}$$

Since for the elliptical distribution  $\mathbb{E}[Q_\eta X|\eta^T X] = 0$ , thus the theorem holds. □

$\mathbb{E}[\text{cov}[Z|Y]] = \text{cov}[Z] - \text{cov}[\mathbb{E}[Z|Y]]$  also could be used to extract information of  $\beta$ s.

# SIR estimation

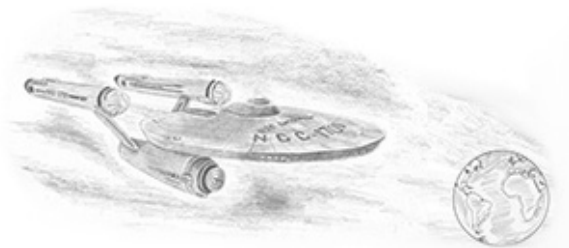
In the case of one-dimensional  $Y$

## Algorithm

- 1 Standardizing  $X$
- 2 Partitioning the whole data into several slices according to the value of  $Y$
- 3 Calculate the slice mean of  $X$  accordingly
- 4 Run principal component analysis on slice means of  $X$
- 5 Locating the most important  $K$ -dimensional subspace for tracking the inverse regression curve  $\mathbb{E}[X|Y]$

## Take home messages

- 1 Don't rely on the models, let the data talk
- 2 The conditional distribution of  $X$  given  $Y$  encodes vital information about dependencies



# Bayesian partitioning for eQTL analysis

## What is eQTL?

- eQTL: expression quantitative trait loci
- To correlate variations in the gene expression with DNA
- cQTL: clinical QTL (traditional GWAS)
- Finding co-localize eQTL and cQTL identifies a list of candidate genes for follow-up studies of the disease

## For imaging-genetic studies

- eQTL  $\Rightarrow$  activations, structural images, connectivities, *etc.*
- To identify a list of genes and imaging traits that correlate with the clinical symptoms.

## Terminologies explained

- cis-acting and trans-acting
  - on the gene or not
- epistatic and pleiotropic effects
  - many to one and one to many

## Some historical comments

- eQTL analysis dates back to a time genome-wide dense sequencing is technically impossible, so it utilizes the LD structure of the genetic markers to identify causal locus.

# Bayesian partitioning (BP) models for eQTL

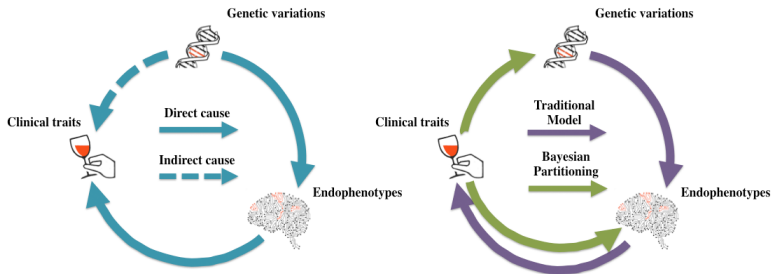
## Highlights

- Integrates eQTL, cQTL and SIR
- Distribution based, indep. of specific interactions
- Accounting for association structures (LD, co-expression)
- Dynamic clustering
- Improved sensitivity for weak couplings

The full model is overwhelmingly sophisticated, so I'll try to capitalize only the key ideas in this talk.



# A peek of causal modeling



**Figure :** (Left) Ground truth causal network (Right) Bayesian causal network used by traditional model (purple) and Bayesian partitioning model (green). Endophenotypes can include gene expression, brain activation, etc.



### Key question for traditional bayesian model

Which models are most consistent with the data **under our assumptions**?

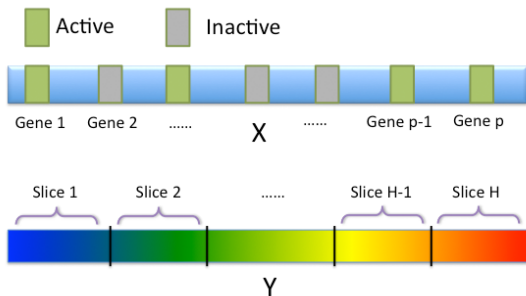
### Key question for Bayesian partition

Which **partition schemes** and **conditional distributions** that are most consistent with the data we observe?

# BP for single quantitative trait

## Basic notations

- $X$ : categorical variables (SNPs),  $X_j \in [1 : K]$
- $Y$ : quantitative trait (gene expression)
- $S(Y)$ : slice membership,  $h \in [1 : H]$
- $\mathcal{A}$ : QTL locus set



## Dirichlet-multinomial model condition on partition

$$X_A | S(Y) = h \sim \text{Multinomial}(1, \theta_A^{(h)})$$

$$\theta_A^{(h)} \sim \text{Dirichlet}\left(\frac{\alpha_0}{K|A|}, \dots, \frac{\alpha_0}{K|A|}\right)$$

## Dynamic partitioning

The slicing prior  $Pr(S(Y)) = \pi_0^{|S|-1} (1 - \pi_0)^{n-|S|}$

Compute  $Pr(X_A | S(Y))$  by integrating out  $\theta_A^{(h)}$

$$Pr(X_A | Y) = \sum_{S(Y) \in \Omega} Pr(X_A | S(Y)) Pr(S(Y))$$

Can be computed in  $O(n^2)$ , draw slicing schemes from  $Pr(S(Y) | X_A, Y)$  via forward - summation - backward - sampling if needed

## Grouping the genes

$I$ : indicator function of active gene set  $\mathcal{A}$

## Saturated NULL model and posterior distribution

$$Pr(X_{\mathcal{A}^c} | X_{\mathcal{A}}, Y) = Pr(X_{\mathcal{A}^c} | X_{\mathcal{A}}) = \frac{Pr_{null}(X)}{Pr_{null}(X_{\mathcal{A}})}$$

$$Pr(I) \sim \text{Bernoulli}(\eta_I, p, |\mathcal{A}|)$$

$$P(I | Y, X) \propto P(X_{\mathcal{A}} | Y) P(X_{\mathcal{A}^c} | X_{\mathcal{A}}) P(I) \propto \frac{Pr(X_{\mathcal{A}} | Y)}{Pr_{null}(X_{\mathcal{A}})} \left( \frac{\eta_I}{1 - \eta_I} \right)^{|\mathcal{A}|}$$

## Bayesian factor and Gibbs sampling

$$BF(\mathcal{A} | Y) = \frac{Pr(X_{\mathcal{A}} | Y)}{Pr_{null}(X_{\mathcal{A}})} = \sum_{S(Y) \in \Omega} BF(X_{\mathcal{A}} | S(Y)) Pr(S(Y))$$

$$BF(X_{\mathcal{A}} | S(Y)) = \frac{Pr(X_{\mathcal{A}} | S(Y))}{Pr_{null}(X_{\mathcal{A}})}$$

$$Pr(I_k = 1 | I_{[-k]}, X, Y) = \frac{\eta_I BF(\mathcal{A}_{[-k]} \cup \{k\} | Y)}{(1 - \eta_I) BF(\mathcal{A}_{[-k]} | Y) + \eta_I BF(\mathcal{A}_{[-k]} \cup \{k\} | Y)}$$

## Multiple conditionally indep. QTL groups

$\mathcal{A}_1, \dots, \mathcal{A}_M$ : conditionally indep. associated gene groups

$$P_m(X_{\mathcal{A}} | S(Y)) = \prod_{m=1}^M P(X_{\mathcal{A}_m} | S(Y))$$

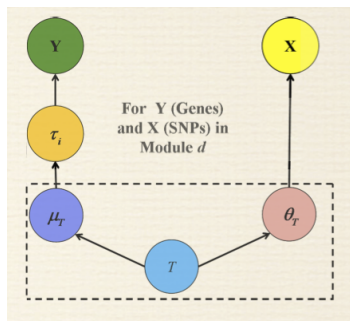
Partition follows Chinese restaurant process

## Modeling block structure of LD

- $L$ : genetic location
- $B, \mathcal{B}_h$ : LD block partition and indicator
- $X_{\mathcal{B}_h} \sim \text{Multinomial}(1, \theta_{\mathcal{B}}^{(h)}), \theta_{\mathcal{B}}^{(h)} \sim \text{Dirichlet}(\frac{\alpha_b}{K^{|\mathcal{B}_h|}}, \dots, \frac{\alpha_b}{K^{|\mathcal{B}_h|}})$
- $P_{blk}(X_{\mathcal{B}_h}), P_{blk}(X|B) = \prod_{h=1}^{|\mathcal{B}|} P_{blk}(X_{\mathcal{B}_h})$
- $P_{blk}(X) = \sum P_{blk}(X|B)P(B), P_{blk}(X_{\mathcal{A}^c} | X_{\mathcal{A}}) = \frac{P_{blk}(X)}{P_{blk}(X_{\mathcal{A}})}$

## Augmented partitioning, gene clustering and multiple modules

- $R, T$ : auxiliary ranking and associated slicing
- $Y_{i,j}$ : gene expressions for subject  $i$ , gene  $j$
- $C_j, \mathcal{G}_c$ : gene cluster membership
- $Y_{i,j} | C_j = c \sim N(\tau_{i,c}, \sigma_c^2)$ ,  $\tau_{i,c} | T_i = t \sim N(\mu_{t,c}, \sigma_c^2 / \kappa_1)$
- $\mu_{t,c} \sim N(0, \sigma_c^2 / \kappa_2)$ ,  $\sigma_c^2 \sim \text{Inv}\chi^2(\nu_0, \sigma_0^2)$

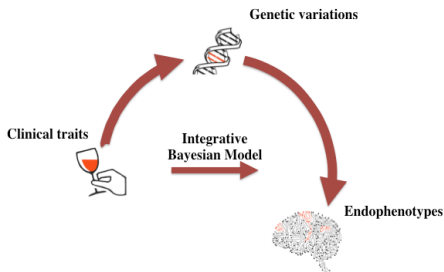




# Comparison with integrative Bayesian model

## Overview of the model in [FC Stingo, 2013, JASA]

- $X \in \mathbb{R}^p$  imaging features,  $Z \in \mathbb{R}^q$  genetic covariates
- $G \in \{1, \dots, K\}$  group indicator
- Latent labels for discriminatory features/covariates
  - $\gamma \in \{0, 1\}^p$  feature label
  - $\delta \in \{0, 1\}^q$  covariate label



## Modeling

- Feature modeling
  - Nondiscriminatory:  $f_0(X_j; \theta_{0j}) \sim N(0, \sigma_{0j}^2)$
  - Discriminatory (group  $k$ ):  $f_k(X_j; \theta_{kj}) \sim N(\mu_{kj}, \sigma_{kj}^2)$
- Covariate effect modeling
  - $\mu_{kj} = \mu_{0k} + \beta_{kj}^T Z$ ,  $\mu_{0k}$  the random effects
  - Sparsity priors on  $\beta_{k(\gamma)}$
- MRF priors for spatial structure

## Comparisons

- Commonalities
  - Sample the latent indicator for feature and covariate
  - Split sample into groups
- Disparities
  - Deterministic VS agnostic grouping
  - Generative VS nongenerative modeling

# Other recent developments

## What we learnt from BP

- SIR is nonparametric, the rest are parametric
- A blend of para. and non-para. ideas might prove useful

## Sliced inverse regression with interaction detection (SIRI)

- Variable selection for active set  $\mathcal{A}$

$$X_{\mathcal{A}}|Y \in \mathcal{S}_h \sim \text{MVN}(\boldsymbol{\mu}_h, \boldsymbol{\Sigma})$$

$$X_{\mathcal{A}^c}|(X_{\mathcal{A}}, Y \in \mathcal{S}_h) \sim \text{MVN}(\alpha + \beta^T X_{\mathcal{A}}, \boldsymbol{\Sigma}_0)$$

- $\boldsymbol{\mu}_h \in \mathbb{V}^q \iff \text{SIR}$
- Likelihood ratio test to compare models
- Forward - addition - backward - deletion

# Concluding remarks

## Limitations

- Where is the p-value
- Difficult to implement and estimate
- Not accounting for the covariate effect
- One dimensional auxiliary ranking

