

### 3.3 Sample path large deviations

In many problems, the interest is actually in rare events that depend on a collection of random variables, or, more generally, on a random process.

Interest often lies in the probability that a path of a random process hits a particular set.

We will briefly address such questions. We start with the simplest example, namely sample path large deviations for random walks.

$X_1, X_2, \dots$  be a sequence of i.i.d random vectors taking values in  $\mathbb{R}^d$ ,  $\Lambda(\lambda) = \mathbb{E} [ e^{\langle \lambda, X_1 \rangle} ] < \infty$  for all  $\lambda \in \mathbb{R}^d$ .

We know that Cramér's theorem gives large deviation principle for the mean  $\frac{1}{n} \sum_{i=1}^n X_i$  (mean vector).

Similarly, the LDP for  $\frac{1}{n} \sum_{i=1}^{\lfloor nt \rfloor} X_i$  can be obtained.

Here, we are seeking the large deviations joint behaviour of a family of random variables indexed by  $t \geq 0$ .

Define

$$Z_n(t) = \frac{1}{n} \sum_{i=1}^{\lfloor nt \rfloor} X_i, \quad 0 \leq t \leq 1,$$

and let  $\mu_n$  be the law of  $Z_n$  in  $L_\infty([0,1]) = L_\infty([0,1]; \mathbb{R}^d)$  (essentially bounded functions

$[0,1] \rightarrow \mathbb{R}^d$ .

Theorem 16 (Mogulskii 1976):

The sequence  $(\mu_n)_{n \geq 1}$  satisfies in  $L_\infty([0,1])$  the LDP with the good rate function

$$I(\phi) = \begin{cases} \int_0^1 \Lambda^*(\dot{\phi}(t)) dt, & \text{if } \phi \in \mathcal{AC}, \phi(0) = 0 \\ \infty & \text{, otherwise,} \end{cases}$$

where  $\mathcal{AC}$  denotes the space of absolutely continuous functions.

Remarks: (1)  $\phi: [0,1] \rightarrow \mathbb{R}^d$  absolutely continuous implies

that  $\phi$  is differentiable almost everywhere (it is an integral of an  $L_1([0,1])$  function).

(2)  $\mu_n$  are supported on the space of functions continuous from the right and having left limits and  $\mathcal{D}_T$  is a subset. Henceforth the LDP holds also when the space is equipped with the  $\|\cdot\|_\infty$ -topology.

(3) Extensions to stochastic processes with jumps at random times exists.



Proof: We give a sketch of the proof and refer to chapter 5 in [DZ98] for further details.

Step 1:  $\tilde{\mu}_n$  law of  $\tilde{Z}_n$  in  $L_\infty([0,1])$

$$\tilde{Z}_n(t) = Z_n(t) + \left(t - \frac{[nt]}{n}\right) X_{[nt]+1}$$

Then  $(\mu_n)_{n \geq 1}$  and  $(\tilde{\mu}_n)_{n \geq 1}$  are exponentially equivalent, i.e., LDP of one ~~implies~~ <sup>implies</sup> the LDP of the other.

Recall that  $(\mu_n), (\nu_n)$  are exponentially equivalent if there is a coupling, that is a probability measure space  $(\Omega, \mathcal{B}_n, \mathbb{P}_n)$  such that  $\mathbb{P}_n^{(1)} = \mu_n$  and  $\mathbb{P}_n^{(2)} = \nu_n$

and  $\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}_n(\Gamma_\delta) = -\infty$ , where,

for  $\delta > 0$ ,  $\Gamma_\delta = \{(x, y) \in E \times E : d(x, y) > \delta\}$ .

Step 2:  $X := \{w : [0,1] \rightarrow \mathbb{R}^d : w(0) = 0\}$ , and equip  $X$  with the topology of pointwise convergence on  $[0,1]$ .

$(\tilde{\mu}_n)_{n \geq 1}$  LDP on  $X$  with rate <sup>good</sup> function  $I$ .

Step 3:  $(\tilde{\mu}_n)_{n \geq 1}$  exponentially tight in  $C_0([0,1])$

(space of all continuous  $w : [0,1] \rightarrow \mathbb{R}^d$ , such that  $w(0) = 0$ , equipped with supremum norm topology).

We outline step 2 and refer to [DZ98] for the remaining ones and the proof of the theorem.

Let  $\mathcal{J}$  denote the collection of all ordered finite subsets of  $[0, 1]$ . For any  $j = \{0 < t_1 < t_2 < \dots < t_{|j|} \leq 1\} \in \mathcal{J}$  and  $f: [0, 1] \rightarrow \mathbb{R}^d$ , let  $p_j(f)$  denote the vector  $(f(t_1), f(t_2), \dots, f(t_{|j|})) \in (\mathbb{R}^d)^{|j|}$ . Then  $(\mu_n \circ p_j^{-1})_{n \geq 1}$  satisfies the LDP in  $(\mathbb{R}^d)^{|j|}$  with the good rate function

$$I_j(z) = \sum_{\ell=1}^{|j|} (t_\ell - t_{\ell-1}) \Lambda^* \left( \frac{z_\ell - z_{\ell-1}}{t_\ell - t_{\ell-1}} \right),$$

where  $z = (z_1, \dots, z_{|j|})$  and  $t_0 = 0, z_0 = 0$ .

To see this define the following random objects:

$$z_n^j = (z_n(t_1), \dots, z_n(t_{|j|})) \quad \text{and}$$

$$Y_n^j = (z_n(t_1), z_n(t_2) - z_n(t_1), \dots, z_n(t_{|j|}) - z_n(t_{|j|-1}))$$

$Y_n^j \mapsto z_n^j$  is  $\pi\pi$  onto itself, continuous and one to one, hence contraction principle applies. LDP for  $(Y_n^j)_{n \geq 1}$  with

rate function

$$\Lambda_j^*(y) := \sum_{\ell=1}^{|j|} (t_\ell - t_{\ell-1}) \Lambda^* \left( \frac{y_\ell}{t_\ell - t_{\ell-1}} \right), \quad y \in (\mathbb{R}^d)^{|j|}$$

It is easy to see that

$$\Lambda_j^*(y) = \sup_{\lambda \in (\mathbb{R}^d)^{|j|}} \left\{ \underbrace{\langle \Lambda, y \rangle}_{\in (\mathbb{R}^d)^{|j|}} - \Lambda_j(\lambda) \right\}, \quad \text{where}$$



$\Lambda_j(\lambda) = \sum_{\ell=1}^{|\mathcal{J}|} (t_\ell - t_{\ell-1}) \Lambda(\lambda e)$ . Thus  $\Lambda_j^*$  is the

Fenchel-Legendre transform of the finite and differentiable function  $\Lambda_j$ . The LDP for  $(Y_n^j)_{n \geq 1}$  follows from the Gärtner-Ellis theorem (Theorem 10), since by independence of  $X_i$ ,  $\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{E} [e^{n \langle \lambda, Y_n^j \rangle}] = \Lambda_j(\lambda)$ .

$(\mu_n \circ p_j^{-1})_{n \geq 1}$  and  $(\tilde{\mu}_n \circ p_j^{-1})_{n \geq 1}$  are exponentially equivalent in  $(\mathbb{R}^d)^{|\mathcal{J}|}$ , thus we also get the following LDP:

For any  $j \in \mathcal{J}$ ,  $(\tilde{\mu}_n \circ p_j^{-1})_{n \geq 1}$  satisfies the LDP in  $(\mathbb{R}^d)^{|\mathcal{J}|}$  with the good rate function  $I_j$  above.

We now turn to the proof of step 2. For  $i, j \in \mathcal{J}$ ,  $i = \{s_1, \dots, s_{|i|}\} \leq j = \{t_1, \dots, t_{|j|}\}$  iff for any  $\ell$ ,  $s_\ell = t_{q(\ell)}$  for some  $q(\ell)$ .

Then, for  $i \leq j \in \mathcal{J}$ , the projection  $p_{ij} : (\mathbb{R}^d)^{|\mathcal{J}|} \rightarrow (\mathbb{R}^d)^{|i|}$  is defined in the natural way.

Let  $X$  denote the projective limit of  $\{Y_j = (\mathbb{R}^d)^{|\mathcal{J}|}\}_{j \in \mathcal{J}}$  with respect to the projections  $p_{ij}$ .

Applying the Dawson-Gärtner theorem (we had no time to discuss that and henceforth refer to [DZ98]) we obtain

the rate function

$$I_X = \sup_{0=t_0 < t_1 < \dots < t_k \leq 1} \sum_{\ell=1}^k (t_\ell - t_{\ell-1}) \Lambda^* \left( \frac{f(t_\ell) - f(t_{\ell-1})}{t_\ell - t_{\ell-1}} \right)$$

where  $f: [0,1] \rightarrow \mathbb{R}^d$ ,  $f(0) = 0$ ,  $f(t) = X_{\xi t}$ .

Note that each point in the projective limit space  $X$ , i.e., each  $x = (x_j)_{j \in \mathbb{J}}$ , may be identified with the map  $f$ .

It remains to show that  $\underline{I}_X = \underline{I}$ .

The convexity of  $\Lambda^*$  implies by Jensen's inequality that  $\underline{I}(\phi) \geq \underline{I}_X(\phi)$ . As for the opposite inequality, first consider  $\phi \in AC$ .

Let  $g(t) := \frac{d\phi(t)}{dt} \in L_1([0,1])$ , and define

$$g^k(t) = k \int_{\lfloor kt \rfloor/k}^{(\lfloor kt \rfloor + 1)/k} g(s) ds, \quad t \in [0,1], \quad k \geq 1, \quad g^k_{(1)} = \int_{\lfloor kt \rfloor/k}^1 g(s) ds$$

$$\underline{I}_X(\phi) \geq \liminf_{k \rightarrow \infty} \sum \frac{1}{k} \Lambda^*(k(\phi(\frac{l}{k}) - \phi(\frac{l-1}{k})))$$

$$= \liminf_{k \rightarrow \infty} \int_0^1 \Lambda^*(g^k(t)) dt$$

By Lebesgue's theorem,  $\lim_{k \rightarrow \infty} g^k(t) = g(t)$  almost everywhere

in  $[0,1]$ . Hence, by Fatou's lemma and the lower semicontinuity of  $\Lambda^*$ ,

$$\begin{aligned} \liminf_{k \rightarrow \infty} \int_0^1 \Lambda^*(g^k(t)) dt &\geq \int_0^1 \liminf_{k \rightarrow \infty} \Lambda^*(g^k(t)) dt \\ &\geq \int_0^1 \Lambda^*(g(t)) dt = \underline{I}(\phi) \end{aligned}$$



## Brownian motion sample path large deviations:

Let  $B_t$ ,  $t \in [0, 1]$  (time horizon), denote a standard Brownian motion in  $\mathbb{R}^d$ . Consider the process

$$B_\varepsilon(t) = \sqrt{\varepsilon} B_t, \text{ and let } \nu_\varepsilon \text{ be the probability}$$

measure induced by  $B_\varepsilon$  on  $C_0([0, 1])$ , the space of all continuous functions  $\phi: [0, 1] \rightarrow \mathbb{R}^d$  such that  $\phi(0) = 0$ , equipped with the supremum norm topology.

$B_\varepsilon$  is a candidate for an LDP similar to the following one (compare with  $Z_n$  above)

$$Y_\varepsilon(t) = \varepsilon \sum_{i=1}^{\lfloor t/\varepsilon \rfloor} X_i, \quad 0 \leq t \leq 1, \quad \text{and } \varepsilon \downarrow 0.$$

First observe the following elementary fact.

Lemma 17: For any  $d \in \mathbb{N}$  and any  $\tau, \varepsilon, \delta > 0$

$$\mathbb{P} \left( \sup_{0 \leq t \leq \tau} |B_\varepsilon(t)| \geq \delta \right) \leq 4d e^{-\delta^2 / 2d\tau\varepsilon}.$$

Proof:

$$\mathbb{P} \left( \sup_{0 \leq t \leq \tau} |B_\varepsilon(t)| \geq \delta \right) = \mathbb{P} \left( \sup_{0 \leq t \leq \tau} |B_t|^2 \geq \varepsilon^{-1} \delta^2 \right)$$

$$\leq d \mathbb{P} \left( \sup_{0 \leq t \leq \tau} (B_t)_1^2 \geq \frac{\delta^2}{d\varepsilon} \right) \text{ because}$$

$$\{x \in \mathbb{R}^d : |x|^2 \geq \alpha\} \subset \bigcup_{i=1}^d \{x \in \mathbb{R}^d : |x_i|^2 \geq \frac{\alpha}{d}\}.$$

The laws of  $B_t$

and  $\sqrt{\varepsilon} B_{t/\varepsilon}$  are identical, we get by time rescaling

$$\mathbb{P}\left(\sup_{0 \leq t \leq \varepsilon} |B_\varepsilon(t)| \geq \delta\right) \leq d \mathbb{P}\left(\|(B_t)_1\| \geq \frac{\delta}{\sqrt{d\varepsilon}}\right).$$

Now define  $B_t^* = (B_t)_1$ , where  $B_t^*$  is a one-dimensional Brownian motion.

$B_t^*$  and  $-B_t^*$  have the same law in  $C_0([0,1])$ ,

$$\mathbb{P}(\|B_t^*\| \geq \eta) \leq 2 \mathbb{P}\left(\sup_{0 \leq t \leq 1} B_t^* \geq \eta\right)$$

$$= 4 \mathbb{P}(B_1^* \geq \eta) \leq 4 e^{-\eta^2/2}, \text{ where the equality}$$

is Désiré André's reflection principle. ■

~~Therefore~~

Thus,  $\|B_\varepsilon\| \xrightarrow{\varepsilon \downarrow 0} 0$  in probability (actually, almost surely) and exponentially fast in  $1/\varepsilon$ .

Theorem 18: (Schilder)

$(\nu_\varepsilon)_{\varepsilon > 0}$  satisfies, in  $C_0([0,1])$ , the LDP with good rate function

$$I_B(\phi) = \begin{cases} \frac{1}{2} \int_0^1 |\dot{\phi}(t)|^2 dt, & \phi \in \mathcal{H}_1, \\ \infty & \text{otherwise.} \end{cases}$$



Remarks: (1)  $H_1 = \{ \phi : \phi(t) = \int_0^t f(s) ds : f \in L_2([0,1]) \}$   
 is the space of all absolutely continuous functions with  
 square integrable derivative equipped with the norm  
 $\|g\|_{H_1} = \left[ \int_0^1 |g(t)|^2 dt \right]^{1/2}$ .

(2) The LDP is for  $\varepsilon \rightarrow 0$ .

Proof:  $\widehat{B}_\varepsilon(t) := B_\varepsilon(\varepsilon \lfloor t/\varepsilon \rfloor)$

is  $Y_\varepsilon$  defined above (i.e.,  $Y_\varepsilon(t) = \varepsilon \sum_{i=1}^{\lfloor t/\varepsilon \rfloor} X_i$ ), for the  
 particular choice of  $X_i$ , which are standard Normal random  
 variables in  $\mathbb{R}^d$  (zero mean and covariance matrix the identity).

Theorem 16 can be easily applied to show that the probability  
 laws of  $\widehat{B}_\varepsilon$  satisfy the LDP in  $L_\infty([0,1])$  with the  
 good rate function  $I(\phi) = \begin{cases} \int_0^1 \Lambda^*(\dot{\phi}(t)) dt, & \text{if } \phi \in \mathcal{A}, \phi(0) = 0 \\ \infty & \text{otherwise.} \end{cases}$

For standard Normal variables considered here, we compute

$$\Lambda(\lambda) = \log \mathbb{E} [ e^{\langle \lambda, X_1 \rangle} ] = \frac{1}{2} |\lambda|^2, \text{ implying}$$

$$\Lambda^*(x) = \sup_{\lambda \in \mathbb{R}^d} \{ \langle \lambda, x \rangle - \frac{1}{2} |\lambda|^2 \} = \frac{1}{2} |x|^2.$$

Hence,  $\mathcal{D}_I = H_1$ , and the rate function  $I$  specialises to  $I_B$ .

For any  $\delta > 0$ ,

$$\mathbb{P}(\|B_\varepsilon - \hat{B}_\varepsilon\| \geq \delta) \leq \left(\frac{1}{\varepsilon}\right) + 1 \mathbb{P}\left(\sup_{0 \leq t \leq \varepsilon} |B_\varepsilon(t)| \geq \delta\right)$$

$$\leq 4d \varepsilon^{-1} (1 + \varepsilon) e^{-\delta^2 / (2d\varepsilon^2)}$$

(follows by semi-homogeneity of increments of the Brownian motion and by Lemma 17).

Consequently,  $\limsup_{\varepsilon \rightarrow 0} \varepsilon \log \mathbb{P}(\|B_\varepsilon - \hat{B}_\varepsilon\| \geq \delta) = -\infty$ ,

and by exponential equivalence it follows that  $(\nu_\varepsilon)_{\varepsilon > 0}$  satisfies the LDP in  $L^\infty([0, 1])$  with the good rate function  $I_B$ . The restriction to  $C_0([0, 1])$  follows since  $B_\varepsilon \in C_0([0, 1])$  with probability one  $\blacksquare$

The preceding results can be applied and extended also to multivariate random walks (multi-index random process) and Brownian sheet.

Real-world applications are quite numerous, among which we just mention asymptotic performance analysis of communication systems (systems of queues in digital networks). The book by Shwartz and Weiss "Large Deviations For Performance Analysis: Queues, Communication and Computing, 1995, Chapman/Hall.

Another application concerns perturbed dynamical systems (Freidlin-Wentzell) which we will study later in chapter 6.



Another area of direct real-world applications are large excursions in  $\mathbb{R}^d$  (applications related to DNA sequence matching, queueing networks, and abrupt change detection in dynamical systems):

$X_i$ , i.i.d.  $\mathbb{R}^d$ -valued random variables with zero mean and finite logarithmic moment generating function  $\Lambda(\lambda)$ .

Let  $b \neq 0$  be a given drift vector in  $\mathbb{R}^d$  and define

$$Z_n = \sum_{i=1}^n (X_i - b)$$

The increments  $Z_n - Z_m$  of the random walk are  $-(n-m)b + \xi_{n-m}$ , where  $\xi_{n-m}$  has zero mean and covariance of the order  $(n-m)$ . Thus, with  $n-m$  large, it becomes unlikely that  $Z_n - Z_m$  is far away from  $-(n-m)b$ . Interest lies in the probability of the rare events  $\{Z_n - Z_m \in A\}$ , for small  $\varepsilon$ , where  $A$  is a closed set that does not intersect the typical ray  $\{ -\alpha b \mid \alpha \geq 0 \}$ .

$$Y_t^\varepsilon = \varepsilon \sum_{i=1}^{\lfloor t/\varepsilon \rfloor} (X_i - b), \quad t \in [0, \infty),$$

rare events  $Y_t^\varepsilon - Y_s^\varepsilon \in A$ .

$T_\varepsilon = \inf \{ t : \exists s \in [0, t) \text{ such that } Y_t^\varepsilon - Y_s^\varepsilon \in A \}$

$\tau_\varepsilon = \sup \{ s \in [0, T_\varepsilon) : Y_{T_\varepsilon}^\varepsilon - Y_s^\varepsilon \in A \} - \varepsilon$ ;  $L_\varepsilon = T_\varepsilon - \tau_\varepsilon$ .

LDP analysis provides estimates and shows  $\varepsilon \log T_\varepsilon$  and  $L_\varepsilon$  converge.